Confirmatory Factor Analysis of the Pittsburgh Sleep Quality Index in Women with Hot Flashes

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Abstract

Objective—Women report poor sleep quality during various stages of the menopause transition and post-menopause, especially those with hot flashes. Sleep measurements vary widely due to the copious instruments available. The Pittsburgh Sleep Quality Index (PSQI) is a frequently used questionnaire that produces a single score for sleep quality. This one-factor structure has not received consistent support in the literature. The goal of this analysis was to determine the best factor structure of the PSQI in women with hot flashes.

Methods—A confirmatory factor analysis was conducted on the PSQI baseline data from three randomized controlled clinical trials enrolling peri- and postmenopausal women with hot flashes (N=849) from the Menopause Strategies: Finding Lasting Answers for Symptoms and Health (MsFLASH) network. Several a priori factor models were compared.

Results—One- and two-factor models did not fit the data. A three-factor model comprising Sleep Efficiency, Perceived Sleep Quality, and Daily Disturbance showed good fit; however, the sleep medication item was dropped due to poor fit and low rates of sleep medication use. The three-factor model was examined in African American (AA) and Caucasian subsamples and found to be similar in both groups; however, two items showed small group differences in strength as indicators.

Conclusions—Sleep quality in midlife women with hot flashes, as measured by the PSQI, appears to comprise three correlated factors. Minor measurement differences detected between groups are of research interest, but do not necessitate different scoring practices. Additional research is needed to further define sleep quality and its associations with health-related outcomes.

MESH Key Words

sleep; sleep quality; measurement; menopause; women; factor analysis
Introduction

Sleep complaints are twice as prevalent in women compared to men, and disturbed sleep contributes to daytime fatigue and poor quality of life. Women report poor sleep quality during various stages of the menopause transition and post-menopause. This is especially true for women experiencing hot flashes, which are prevalent in as many of 50% of women during the late menopause transition and 48% during early postmenopause.

Selecting a measure of sleep to use in research studies largely depends on the research question(s). Current self-report sleep measures typically address single or combinations of sleep complaints, such as: (a) sleep duration or total sleep time; (b) sleep latency (minutes to fall asleep); (c) number and duration of nighttime awakenings; (d) quality of perceived sleep; (e) sleep efficiency (amount of time in bed spent asleep vs amount of time spent in bed); (f) sleep medication use, and (g) daytime dysfunction (daytime sleepiness, inability to function during the day, feeling tired or fatigued).

Instruments address either a single or combination of the above aspects and are often specific to the overarching sleep problem being addressed. Recommendations have been made for the standardization of sleep measures in order to facilitate meta-analyses in research in patients with cancer and insomnia, but a variety of instruments continue to be used.

The Pittsburgh Sleep Quality Index (PSQI) is one of the most frequently used questionnaires for the assessment of self-reported sleep quality in studies of men and women with and without chronic illnesses. The PSQI was created to capture sleep quality in a sample of psychiatric patients. The original intent of the questionnaire was to: (a) create a reliable and valid standardized measure of sleep quality; (b) provide a tool to distinguish good versus poor sleep quality; (c) create a measure of sleep quality that was user friendly for patients, clinicians, and researchers to interpret; and (d) provide a clinical tool that assessed a list of sleep disturbances that impact sleep quality for psychiatric populations. The PSQI contains 7 individual sub-scales and a single factor global score of > 5 representing poor overall sleep quality. The global score of the PSQI has been deemed a simple measure for use in clinical and research to identify good versus poor sleep. The PSQI single-factor based score for sleep quality has consistently reported acceptable reliability and validity in various populations and has been translated into several different languages.

The single factor structure of the PSQI has been analyzed to understand if the global sleep quality score derived from all of the individual 7- subscale scores is the best representation of sleep quality. This research, conducted in various adult samples with (depression, breast cancer, post-renal transplantation, rheumatoid arthritis) and without chronic illnesses (community dwelling English and Spanish men and women, Nigerian students, non-depressed men and women), has yielded inconsistent results. Of seven studies, four found that a model comprising three factors (e.g. Sleep Efficiency, Perceived Sleep Quality, and Daytime Functioning) was a better fit to the data than a 1-factor global score for sleep quality. One study found a 1-factor model (Sleep Quality) and two studies reported 2-factor models (Sleep Efficiency, Perceived Sleep Quality) were most appropriate for the populations being studied. The Otte et al. study also found racial differences in a nested-model comparison of African American and Caucasian breast cancer survivors of Sleep...
Quality and Sleep Latency as well as Sleep Efficiency and Sleep Quality. The findings from these studies raise questions about whether it is appropriate to use a single global score to differentiate good from poor sleep quality across clinical populations. There were several limitations to these previous studies that make it difficult to generalize the findings of sleep quality to other populations, such as women with hot flashes. These include sample heterogeneity (mixed gender, wide age ranges and specific focus on chronic illnesses) and varied sample sizes from 107 to 1174. There have been no multi-group analyses using factor modeling to determine if mixed sample groups limit the factor findings. In addition, findings have been limited to research studies and it is unclear how findings can be translated into clinical practice.

Since sleep is a significant problem in women during the menopause transition, it would be informative to evaluate whether the single factor measure from the PSQI compared to the previously reported Otte et al. 2-factor model and Cole et al. 3-factor models best reflects good versus poor sleep quality in a large sample of relatively healthy midlife women in late perimenopause and early post-menopause with hot flashes. Additional nested model comparisons between African American and Caucasian women can identify potential racial differences which have been reported in a previous study of women with breast cancer. The goal of this analysis is to provide recommendations on the best factor structure of the PSQI to analyze sleep quality in research trials of women with hot flashes.

**METHODS**

**Procedures**

Baseline data from the PSQI were pooled from three samples of midlife women participating in Menopause Strategies: Finding Lasting Answers for Symptoms and Health (MsFLASH) controlled randomized trials. Study 1 was a randomized control trial of escitalopram versus placebo for the treatment of hot flashes. Study 2 was a 3 by 2 factorial design study that yielded 3 comparison groups: (a) yoga versus usual activity; (b) exercise versus usual activity; and (c) omega 3 fatty acids versus placebo pill for hot flashes. Study 3 was a randomized control trial of low-dose estradiol versus placebo and venlafaxine XR versus placebo for menopausal symptoms. Data from a total of 899 midlife women from these 3 studies were evaluated (study 1 n = 205, study 2 n = 355, study 3 n = 339). Descriptions of the procedures used in the MsFLASH trials have been published elsewhere.

**Setting**

Participants were recruited from five MsFLASH network sites (Seattle, Boston, Oakland, Indianapolis, and Philadelphia). Participants were recruited from July 2009 to October 2012 through targeted mailings to midlife aged-women using purchased mailing lists.

**Measures**

Sample characteristics were collected using an adapted multi-item questionnaire that included items from the PSQI and also measured race, age, marital status, educational and income level, employment status, height, and weight.
Sleep was assessed using the PSQI, a 19-item scale that provides 7 component scores (ranges 0–3): sleep quality, sleep latency, sleep duration, habitual sleep efficiency, sleep disturbance, use of sleep medications, and daytime dysfunction. The 7 component scores are combined to produce a global sleep quality index score. The global scores range from 0–21, with scores above 5 reflecting poor global sleep quality and scores above 8 reflecting poor sleep and high daytime fatigue burden.

**Statistical Analysis**

In order to determine the best factor model and scoring method for the PSQI among women with hot flashes, a confirmatory factor analysis using LISREL 8.8 was conducted. Because the single global score of the PSQI is frequently used in the literature, we examined the fit of a 1-factor model. Because recent studies (Cole et al. and Otte et al.) found support both for 2-factor and 3-factor models, we examined the fit of these models in the present sample. Given that parsimonious models are preferred to more complex ones, we examined these models in order of increasing complexity.

The subscale scores from the 7 PSQI components represent ordinal rather than continuous data and should not be analyzed using methods that assume they are continuous variables with metric properties. Instead, PRELIS 2.8 was used to estimate the polychoric correlations among the subscale scores. Polychoric correlations are used to describe the associations between observed ordinal variables (i.e., PSQI subscale scores) that represent underlying phenomena (i.e., sleep characteristics) that are normally distributed and continuous. These polychoric correlations were then used to create an asymptotic covariance matrix (similar to a correlation matrix) that represents all of the inter-relationships among the PSQI subscales. It is this matrix of associations among the PSQI subscales that was used to examine the various factor models (using the weighted least squares approach).

The appropriateness of a factor model is evaluated based on how well the theoretical model fits the observed matrix of associations. Several indices have been developed to describe model fit, with each index using different criteria. We chose three of the most commonly-used fit indices: (a) chi-square ($\chi^2$); (b) the Root Mean Square Error of Approximation (RMSEA) and its 90% confidence interval; and (c) the Comparative Fit Index (CFI). The $\chi^2$ is an index of absolute fit between the hypothesized and observed covariance matrices. A non-significant $\chi^2$ is evidence of acceptable fit. The RMSEA is a parsimonious index, meaning that the complexity of the hypothesized model is taken into consideration in evaluating its fit with the observed covariance matrix. RMSEA values ≤ .06 indicate acceptable fit. The CFI is an incremental fit index, with values ≥ .95 indicating acceptable fit.

Once the best-fitting model has been determined in the overall sample, its consistency across subsamples can be examined using nested-model comparisons. Such comparisons involve statistically comparing different parameters of the model to find which ones are significantly different in one group versus the other.
RESULTS

Sample characteristics

The majority of the women were non-Hispanic (97.2%), Caucasian (59.0%), partnered (63.2%), employed (69.7%), and highly-educated (79.6% with some college). A substantial portion of the sample was African American (33.8%), with smaller proportions of American Indian (1.2%), Asian/Pacific Islander (2.2%), and women who self-identified as “Other” (2.8%). The mean age was 54.47 years (SD = 3.83). The mean global PSQI score was 7.82 (SD = 3.47). The majority of women (72.0%) had global scores above the cutoff score of 5, suggesting poor sleep quality, and a sizeable portion of those women (39.8%) had scores >8 indicating poor sleep quality and high fatigue burden.

Responses to the PSQI were scored into the seven subscales and 94.4% of the sample had complete baseline data on all component scores. The means, standard deviations, and correlations among the subscales are shown in Table 1. The number of missing values for each subscale ranged from 7 to 17, with a total of 50 participants missing data on at least one subscale. None of the cases with missing values had adequate information to allow for imputation. Therefore, these 50 cases were excluded from the confirmatory factor analysis, resulting in a final sample size of 849. We compared women with complete data to women with missing data and found that there were no differences on age (M = 54.48 & 54.16 for complete and missing data respectively). There was a small, but significant relationship between education and missing data points; women with less education were more likely to have missing data, $r_s = -0.078$, $p = 0.019$. There also was a relationship between race/ethnicity and missing data, $\chi^2(5) = 19.04$, $p = .002$; specifically, proportionately more African American women had missing data compared to Caucasian women (7.89% vs. 2.19%, respectively). Similarly, never married women had proportionately more missing data than women who were married or divorced/separated (8.47% vs. 3.40% & 2.16%, respectively, $\chi^2(4) = 9.51$, $p = .049$). This could be attributed to women without bed partners not providing a response for the bed partner related item.

Confirmatory factor analysis

The sample sizes, means, standard deviations, and correlations among the seven PSQI subscales are shown in Table 1. Univariate skew and kurtosis statistics were examined for each component score and none of the indicators showed excessive skew or kurtosis. Hence, none of the data were transformed. The fit indices for the tested models are shown in Table 2. The 1-factor model (Model 1), consistent with the global score of the PSQI, did not show acceptable fit on any of the indices. The 2-factor model (Sleep Efficiency, Perceived Sleep Quality) (Model 2) also did not show acceptable fit on any index. A modified 2-factor model (Model 3) showed better fit, with the CFI indicating acceptable fit. The 3-factor model (Sleep Efficiency, Perceived Sleep Quality; Daily Disturbances) (Model 4) showed the best fit, with acceptable values on the RMSEA and CFI.

The 3-factor model (Model 4) is presented in Figure 1 with its standardized coefficients. As the model shows, the loading from Perceived Sleep Quality to the sleep medication use subscale is small ($\beta = .15$), which leaves a large amount of unexplained variability in this...
item ($\delta = .98$). The small loading of the sleep medication use item is consistent with its weak correlation with the other subscales of the PSQI (Table 1). As the correlation table shows, even the strongest association of sleep medication use with sleep latency was small ($r = .15$). This is likely due to a restriction of range, as 68% of the sample reported using no sleep medication at all.

Given the low rate of sleep medication use and the resultant weak correlations between it and the other PSQI subscales, we examined an alternative 4-factor model where Sleep Medication Use was modeled as a separate factor along with Perceived Sleep Quality, Sleep Efficiency, and Daily Disturbance. This model showed good fit to the data (Model 5 in Table 2). However, the Sleep Medication Use factor did not correlate strongly with any of the other factors ($\psi = .03-.18$), which raises the question of utility of this item in assessing sleep quality among women with hot flashes.

Consequently, we examined the fit of an alternative 3-factor model (Sleep Efficiency, Perceived Sleep Quality, and Daily Disturbance), excluding the sleep medication use item, and this model showed good fit to the data across all indices, including a non-significant $\chi^2$ (Table 2; Model 6). Therefore, we concluded that this 3-factor conceptualization represented the best model of sleep quality among women with hot flashes (see Figure 2).

To examine whether the factor structure was consistent between African American and Caucasian subsamples, we conducted several two-group comparisons. Although the sample consisted of women from several racial/ethnic groups, only the African American (n = 304) and Caucasian (n = 530) subsamples had sufficient numbers to allow for statistical comparisons of the factor models. First, a two-group, 3-factor model was examined with all parameters freed to vary between groups. This model showed good fit to the data across all indices (Table 2; Model 6a). We constrained the factor loadings to be equal across both groups, and the resulting 3-factor model showed significantly worse fit to the data, $\Delta\chi^2 (df = 6) = 23.81, p = 0.0005$, indicating that the factor loadings differed between African American and Caucasian women.

Subsequently, we conducted a series of nested-model comparisons to determine which factor loadings onto the PSQI subscales differed between the two groups (Table 3). The loading values for the Perceived Sleep Quality factor onto the subjective sleep quality subscale differed significantly between groups, with standardized loadings of .72 and .87 for Caucasian and African American women, respectively. This suggests that subjective sleep quality subscale is a slightly stronger indicator of Perceived Sleep Quality factor in African American women compared to Caucasian women. In addition, the loading values for the Sleep Efficiency factor onto the habitual sleep efficiency subscale differed significantly between groups, with standardized loadings of .73 and .44 for Caucasian and African American women, respectively. This indicates that among menopausal women, habitual sleep efficiency subscale is a stronger indicator of Sleep Efficiency factor for Caucasian women compared to African American women. Although we were able to detect two differences in the factor loadings of African American compared to Caucasian women, these results do not require different PSQI scoring procedures for the two groups. For both groups of women, subjective sleep quality and habitual sleep efficiency are reliable, positive indicators of their respective constructs. These loading differences may indicate subtle
cultural differences in expectations or habitual practices with regard to sleep and sleep-related behaviors.

However, these differences are mainly of interest for research purposes and do not necessitate different scoring procedures for African American versus Caucasian women.

Next, we constrained the correlations among the three factors to be equal between groups to test if the factor intercorrelations differed in African American and Caucasian women. The resulting model showed no significant difference in fit, $\Delta \chi^2 (df = 3) = 5.34, p = 0.1485$, indicating that the correlations among Sleep Efficiency, Perceived Sleep Quality, and Daily Disturbance were the same in African American and Caucasian women. This result suggests that the structure of sleep quality is similar in African American and Caucasian women with hot flashes and that the pattern of relationships among the factors is the same for both groups. Taken together, these results suggest that the three-factor model is the best conceptualization of the PSQI among women with hot flashes and that there are no major racial differences.

DISCUSSION

The results of this study showed that sleep quality, as measured by the PSQI, in midlife women in the late menopause transition and early postmenopause with daily hot flashes is multifaceted with three correlated factors: Sleep Efficiency, Perceived Sleep Quality, and Daily Disturbance. This 3-factor model fit well in both African American and Caucasian subsamples and the correlations among the factors were equivalent across the two groups. This finding suggests that from a practical perspective, the structure of sleep quality, at least as measured by the PSQI, is consistent in African American and Caucasian women.

Overall, results of these analyses in midlife women with hot flashes are consistent with previous reports analyzing a three-, two- and one factor models of sleep quality. The nested model comparisons by race show no major differences as with the previous Otte et al. study. The model results are similar to previous factor analyses that also found that a 3-factor model (Sleep Efficiency, Perceived Sleep Quality, Daily Disturbances) best fit the data in depressed and non-depressed adults, 520 Nigerian university students, 135 post-renal transplant patients, and 3,667 community-dwelling English and Spanish speaking Hispanic and non-Hispanic adults. Conversely, the traditional 1-factor (Sleep Quality) and 2-factor models (Sleep Efficiency; Perceived Sleep Quality) consisted of samples of 197 Chinese women with breast cancer, 417 depressed and non-depressed older adults, 107 patients with rheumatoid arthritis, and 1174 non-depressed breast cancer survivors. Two of the prior studies also produced a better model fit by removing the sleep medication item during the analysis. The findings from these various studies show that the PSQI factor structure differ among studies that have different mixed genders, race, and chronic illnesses. Differences in factor structure can also be attributed to the wide range of sample sizes among the studies, which can impact factor structure. The result of this variability in samples and factor structures suggests future research studies should consider how the PSQI should best be analyzed in light of these findings. The results also suggest that when
evaluating intervention efficacy using the PSQI using different factor structure can delineate improvements in more specific areas such as sleep efficiency or perceived sleep quality.

Although we found that the factor structure and factor inter-correlations were consistent between African American and Caucasian women, nested model comparisons did reveal two measurement differences. First, we found that the subjective sleep quality subscale was a stronger indicator of Perceived Sleep Quality factor in African American women compared to Caucasian women. This difference is important from a measurement and research standpoint. However, it is unlikely to be of any practical or clinical significance. For both African American and Caucasian women, subjective sleep quality is the best indicator of Perceived Sleep Quality. Second, we found that assessing the Sleep Efficiency factor differed significantly in African American versus Caucasian women. Specifically, we found that Sleep Efficiency among African American women was primarily captured by the sleep duration subscore. In contrast, the habitual sleep efficiency subscale was a stronger indicator of Sleep Efficiency in Caucasian versus African American women. This could mean that African Americans with poor sleep efficiency have shorter sleep duration. Again, this difference is important from a research and measurement perspective, but is unlikely to matter in the clinical assessment of sleep quality.

Our results suggest that the sleep medication item may not be a good indicator of sleep quality among women participating in clinical trials for hot flashes, likely due to low use of sleep medication in this population. This is a consistent finding across the multifactor model analyses. In the present sample, 68% of women reported not using any sleep medications at all. Sleep medication use may be an indicator of sleep quality among older women or those being treated for illnesses (e.g., breast cancer), as they may be more likely to take medications to improve different aspects of quality of life. Further research is needed to determine what factors drive the use of medications to help with sleep among women. Although it is clinically important to query patients regarding sleep medication use, it is recommended that the item be removed from the PSQI sub-scale scoring until further validity and reliability testing is performed.

The results from this study and the previous factor analyses highlight the fact that the definition of sleep quality varies among researchers, clinicians and patients resulting in a poorly defined concept. The concept of perceived sleep quality tends to have a single meaning based on subjective criteria as a way to interpret the overall perception of an individual’s sleep. A qualitative study of adults with and without insomnia found that perceived sleep quality is a multifaceted concept that includes: (a) being tired upon waking that lasts throughout the day; (b) the feeling of being rested upon waking; and (c) number of awakenings during the night of sleep. It could be suggested that assessing patient’s sleep quality requires appraisal of sleep as a continuous assessment of multiple factors and cannot be distilled into a single factor concept as in the PSQI. However, the data findings in this analysis are limited to women in late menopause transition and early postmenopausal experiencing hot flashes limiting generalizability to all midlife women of this age range and menopausal status.
CONCLUSION

Because of the wide range of research instruments used to measure the different aspects of self-reported sleep, it is challenging to aggregate findings across studies to compare the incidence and prevalence of sleep complaints. This is especially true in women who continue to remain underrepresented in sleep research. The PSQI, as a measure of sleep quality, is sensitive to race, ethnicity, chronic illness, and gender, making cross-study comparisons of sleep a continued challenge.

Based on the results of this study, alternative PSQI scoring could be considered and further evaluated using the 3-factor model for self-reported sleep in studies of midlife women with hot flashes. Revisions would consist of recalculations of subscale loadings into Sleep Efficiency (sleep disturbance and habitual sleep efficiency subscale), Perceived Sleep Quality (subjective sleep quality and sleep latency subscales, and Daily Disturbance (sleep disturbance and daytime dysfunction subscale). There is also the issue of how sleep quality is defined and the problem with a single score reflecting a complex concept. Further work is needed to address this issue for measuring sleep quality. Improving understanding of the larger concept of sleep quality can further research efforts to facilitate better subjective measurement that can be used for intervention testing that can improve sleep for women during the menopause transition.

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Conflicts of interest and Source of Funding:

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REFERENCES


Figure 1.
3-Factor Model of Sleep Quality Measured with the Pittsburgh Sleep Quality Index in Midlife Women\textsuperscript{15}

Note: All coefficients are standardized and significant at $p < .05$, unless otherwise noted. Ovals = latent variables. Rectangles = measured variables. Single-headed arrows = factor loadings. Double-headed arrows = correlations; Italicized coefficients are error terms, representing the proportion of variation in the variable not accounted for by the model.
Figure 2.
Alternative 3-Factor Model of Sleep Quality Measured with the Pittsburgh Sleep Quality Index in Midlife Women without Sleep Medication Use Item
Note: All coefficients are standardized and significant at p < .05, unless otherwise noted. Ovals = latent variables. Rectangles = measured variables. Single-headed arrows = factor loadings. Double-headed arrows = correlations; Italicized coefficients are error terms, representing the proportion of variation in the variable not accounted for by the model; † = Parameters were found to differ significantly between African American and Caucasian subsamples. A = parameter estimate for African American subsample; C = parameter estimate for Caucasian subsample.
Table 1

Complete Data Sample Sizes, Means, Standard Deviations, and Correlations of 7 Pittsburgh Sleep Quality Index Component Scores

<table>
<thead>
<tr>
<th>Component</th>
<th>N</th>
<th>Mean</th>
<th>SD</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
<th>6</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. subjective sleep quality</td>
<td>890</td>
<td>1.62</td>
<td>0.74</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
</tr>
<tr>
<td>2. sleep latency</td>
<td>885</td>
<td>1.33</td>
<td>1.01</td>
<td>39**</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
</tr>
<tr>
<td>3. sleep duration</td>
<td>890</td>
<td>0.83</td>
<td>0.93</td>
<td>46**</td>
<td>32**</td>
<td>--</td>
<td>--</td>
<td>--</td>
<td>--</td>
</tr>
<tr>
<td>4. habitual sleep efficiency</td>
<td>888</td>
<td>0.85</td>
<td>1.06</td>
<td>30**</td>
<td>27**</td>
<td>55**</td>
<td>--</td>
<td>--</td>
<td>--</td>
</tr>
<tr>
<td>5. sleep disturbances</td>
<td>892</td>
<td>1.64</td>
<td>0.57</td>
<td>28**</td>
<td>17**</td>
<td>15**</td>
<td>16**</td>
<td>--</td>
<td>--</td>
</tr>
<tr>
<td>6. sleep medication use</td>
<td>884</td>
<td>0.61</td>
<td>1.03</td>
<td>07*</td>
<td>15**</td>
<td>01</td>
<td>03</td>
<td>10**</td>
<td>--</td>
</tr>
<tr>
<td>7. daytime dysfunction</td>
<td>882</td>
<td>0.97</td>
<td>0.75</td>
<td>31**</td>
<td>16**</td>
<td>19**</td>
<td>08**</td>
<td>29**</td>
<td>06</td>
</tr>
</tbody>
</table>

Score ranges 0–3.

Note: Ns for correlations vary between 875 and 890 due to missing data.

* $p < .05$

** $p < .01$
Table 2

Fit Indices for Different Factor Models of the Pittsburgh Sleep Quality Index

<table>
<thead>
<tr>
<th>Model</th>
<th>Factors</th>
<th>$\chi^2$</th>
<th>df</th>
<th>p</th>
<th>RMSEA</th>
<th>90% CI: RMSEA</th>
<th>CFI</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1-factor (Overall Sleep Quality)</td>
<td>107.84</td>
<td>14</td>
<td>&lt;0.00001</td>
<td>0.089</td>
<td>0.074–0.100</td>
<td>0.90</td>
</tr>
<tr>
<td>2</td>
<td>2-factor (Sleep Efficiency, Perceived Sleep Quality)</td>
<td>62.71</td>
<td>13</td>
<td>&lt;0.00001</td>
<td>0.067</td>
<td>0.051–0.084</td>
<td>0.94</td>
</tr>
<tr>
<td>3</td>
<td>2-factor (Sleep Efficiency, Perceived Sleep Quality)</td>
<td>41.97</td>
<td>8</td>
<td>&lt;0.00001</td>
<td>0.071</td>
<td>0.051–0.093</td>
<td>0.96</td>
</tr>
<tr>
<td>4</td>
<td>3-factor (Sleep Efficiency, Perceived Sleep Quality; Daily Disturbances)</td>
<td>31.92</td>
<td>11</td>
<td>0.00079</td>
<td>0.047</td>
<td>0.029–0.067</td>
<td>0.98</td>
</tr>
<tr>
<td>5</td>
<td>4-factor (Sleep Efficiency, Perceived Sleep Quality; Daily Disturbances, Sleep Medication Use)</td>
<td>25.31</td>
<td>9</td>
<td>0.00265</td>
<td>0.046</td>
<td>0.025–0.068</td>
<td>0.98</td>
</tr>
<tr>
<td>6</td>
<td>3-factor (Sleep Efficiency, Perceived Sleep Quality; Daily Disturbances; without sleep medication use indicator)</td>
<td>12.27</td>
<td>6</td>
<td>0.05600</td>
<td>0.035</td>
<td>0.000–0.063</td>
<td>0.99</td>
</tr>
<tr>
<td>6a</td>
<td>3-factor (Sleep Efficiency, Perceived Sleep Quality; Daily Disturbances; without sleep medication use indicator (2 group))</td>
<td>23.08</td>
<td>12</td>
<td>0.02709</td>
<td>0.048</td>
<td>0.016–0.078</td>
<td>0.99</td>
</tr>
</tbody>
</table>
### Table 3
Chi Square Tests Comparing Factor Loadings of the Pittsburgh Sleep Quality Index in African America and Caucasian Women

<table>
<thead>
<tr>
<th>Factor</th>
<th>Indicator</th>
<th>$\lambda^A$</th>
<th>$\lambda^C$</th>
<th>$\Delta \chi^2$</th>
<th>df</th>
<th>p</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sleep Efficiency</td>
<td>sleep duration</td>
<td>0.95</td>
<td>0.86</td>
<td>0.81</td>
<td>1</td>
<td>0.368</td>
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<tr>
<td></td>
<td>habitual sleep efficiency</td>
<td>0.44</td>
<td>0.73</td>
<td>14.66</td>
<td>1</td>
<td>&lt;0.001</td>
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<tr>
<td>Perceived Sleep Quality</td>
<td>subjective sleep quality</td>
<td>0.87</td>
<td>0.72</td>
<td>4.21</td>
<td>1</td>
<td>0.040</td>
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<tr>
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<td>sleep latency</td>
<td>0.51</td>
<td>0.51</td>
<td>0.00</td>
<td>1</td>
<td>1.000</td>
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<tr>
<td>Daily Disturbance</td>
<td>sleep disturbances</td>
<td>0.47</td>
<td>0.54</td>
<td>0.44</td>
<td>1</td>
<td>0.507</td>
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<tr>
<td></td>
<td>daytime disturbance</td>
<td>0.54</td>
<td>0.65</td>
<td>1.11</td>
<td>1</td>
<td>0.292</td>
</tr>
</tbody>
</table>

Note: $\lambda^A$ = factor loading in African American subsample; $\lambda^C$ = factor loading in Caucasian subsample